

Main results on distributions used in Dynare. We use the parametrization used in Distrubtions.jl

## 1 Inverse gamma

- AKA inverse gamma type 2
- shape:  $\alpha$
- scale:  $\theta$

$$\begin{aligned}\alpha &= \frac{\nu}{2} \\ \theta &= \frac{s}{2} \\ X \sim IG_2(\alpha, \theta) &\Leftrightarrow Z = X^{-1} \sim G(\alpha, \theta) \\ \mathbb{E}(X) &= \frac{\theta}{\alpha - 1} \\ \mathbb{V}ar(X) &= \frac{1}{\alpha - 2} [\mathbb{E}(X)]^2 \text{ for } \alpha > 2 \\ \alpha &= 2 + \frac{[\mathbb{E}(X)]^2}{\mathbb{V}ar(X)} \\ \theta &= (\alpha - 1) \mathbb{E}(X) \\ f_{IG_2}(x, \alpha, \theta) &= \frac{\theta^\alpha}{\Gamma(\alpha)} x^{-(\alpha+1)} e^{-\frac{\theta}{x}}\end{aligned}$$

## 2 Inverse gamma type I

$$\begin{aligned}X \sim IG_2(\alpha, \theta) &\Leftrightarrow Y = \sqrt{X} \sim IG_1(\alpha, \theta) \\ &\Leftrightarrow Z = X^{-1} \sim G(\alpha, \theta) \\ \mathbb{E}(Y) &= \sqrt{\theta} \frac{\Gamma(\alpha - \frac{1}{2})}{\Gamma(\alpha)} \text{ for } \alpha > \frac{1}{2} \\ \mathbb{V}ar(Y) &= \frac{\theta}{\alpha - 1} - [\mathbb{E}(Y)]^2 \text{ for } \alpha > 1 \\ \text{mode } Y &= \sqrt{\frac{\theta}{\alpha + \frac{1}{2}}} \\ f_{IG_1}(y, \alpha, \theta) &= f_{IG_2}(y^2, \alpha, \theta) |2y| \\ &= 2 \frac{\theta^\alpha}{\Gamma(\alpha)} y^{-(2\alpha+1)} e^{-\frac{\theta}{y^2}}\end{aligned}$$

$\alpha$  solves

$$(\alpha - 1) (\mathbb{V}ar(Y) + [\mathbb{E}(Y)]^2) - [\mathbb{E}(Y)]^2 \frac{\Gamma(\alpha)}{\Gamma(\alpha - \frac{1}{2})} = 0$$

and  $\theta = (\alpha - 1)(\mathbb{V}ar(Y) + [\mathbb{E}(Y)]^2)$

## Appendices

### Bauwens et (1999)

$$\begin{aligned} X &\sim IG_2(s, \nu) \Leftrightarrow Y = \sqrt{X} \sim IG_1(s, \nu) \\ &\Leftrightarrow Z = X^{-1} \sim G\left(\frac{\nu}{2}, \frac{2}{s}\right) \\ \mathbb{E}(X) &= \frac{s}{\nu - 2} \\ \mathbb{V}ar(X) &= \frac{2}{\nu - 4} [\mathbb{E}(X)]^2 \text{ for } \nu > 4 \\ \mathbb{E}(Y) &= \sqrt{\frac{s}{2}} \frac{\Gamma\left(\frac{\nu-1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)} \text{ for } \nu > 1 \\ \mathbb{V}ar(Y) &= \frac{s}{\nu - 2} - [\mathbb{E}(Y)]^2 \text{ for } \nu > 2 \\ f_{IG_2}\left(x \middle| \frac{\nu}{2}, \frac{2}{s}\right) &= \frac{1}{\Gamma\left(\frac{\nu}{2}\right) \left(\frac{2}{s}\right)^{\frac{\nu}{2}}} x^{-\frac{1}{2}(\nu+2)} e^{-\frac{s}{2x}} \\ f_{IG_2}(x|\alpha, \theta) &= \frac{1}{\Gamma(\alpha)\theta^\alpha} x^{-(\alpha+1)} e^{-\frac{1}{sx}} \end{aligned}$$

### Inverse gamma type I

Mode derivation

$$\begin{aligned} \max_x y &= 2 \frac{\theta^\alpha}{\Gamma(\alpha)} x^{-(2\alpha+1)} e^{-\frac{\theta}{x^2}} \\ \frac{dy}{dx} &= -[(2\alpha + 1)x^{-1} - 2\theta x^{-3}] y \\ &= 0 \\ x^\star &= \sqrt{\frac{\theta}{\alpha + \frac{1}{2}}} \end{aligned}$$

Obtaining  $\alpha$  and  $\theta$  from mean and variance

$$\begin{aligned} \theta &= [\mathbb{E}(Y)]^2 \left( \frac{\Gamma(\alpha)}{\Gamma(\alpha - \frac{1}{2})} \right)^2 \\ \theta &= (\alpha - 1) (\mathbb{V}ar(Y) + [\mathbb{E}(Y)]^2) \end{aligned}$$

Solve numerically

$$(\alpha - 1) (\mathbb{V}ar(Y) + [\mathbb{E}(Y)]^2) - [\mathbb{E}(Y)]^2 \left( \frac{\Gamma(\alpha)}{\Gamma(\alpha - \frac{1}{2})} \right)^2 = 0$$

or taking the logarithm for numerical stabilitye:

$$\ln(\alpha - 1) + \ln (\mathbb{V}ar(Y) + [\mathbb{E}(Y)]^2) - \ln [\mathbb{E}(Y)]^2 - 2 * \ln \left( \frac{\Gamma(\alpha)}{\Gamma(\alpha - \frac{1}{2})} \right) = 0$$